

Chapter 21

The Cox-Ingersoll-Ross Process

In the following we list some named models for the short rate process. We follow the presentation of [4].

The Vasicek Model

The Vasicek model is a mean reverting Ornstein-Uhlenbeck process. It is given by the SDE

$$dr_t = a(b - r_t)dt + \sigma dB_t \quad (21.1)$$

with non-negative constants a, b and σ and initial value $r_0 > 0$. Its solution is given by (compare lecture 14)

$$r_t = r_0 e^{-at} + b(1 - e^{-at}) + \sigma e^{-at} \int_0^t e^{as} dB_s \quad (21.2)$$

Thus r_t has a Gaussian distribution function with mean $r_0 e^{-at} + b(1 - e^{-at})$ and variance $\sigma^2 e^{-2at} \int_0^t e^{2as} ds = \frac{\sigma^2}{2a}(1 - e^{-2at})$. In particular, the probability for negative values of r_t is non zero which makes the model not too realistic. Its consideration is merely due to its simplicity.

The Hull-White Model

The Hull-White model in its simplest form is a Vasicek model with time dependent, deterministic coefficients. It is given by the SDE

$$dr_t = (\alpha(t) - \beta(t)r_t)dt + \sigma(t) dB_t \quad (21.3)$$

with some initial value $r_0 > 0$. The solution is $(b(t) = \int_0^t \beta(s) ds)$

$$r_t = e^{-b(t)} \left(r_0 + \int_0^t e^{b(s)} \alpha(s) ds + \int_0^t e^{b(s)} \sigma(s) dB_s \right) \quad (21.4)$$

which is again a Gaussian process. The model is popular with practitioners. In its more general form it includes a term r_t^γ in the volatility, in which case it generalizes the Cox-Ingersoll-Ross model discussed next.

The Cox-Ingersoll-Ross Model

The following material is taken from section 9.6 of [4]. In the Vasicek and Hull-White model the probability for a negative rate r_t is non zero. The Cox-Ingersoll-Ross model provides an SDE for r_t whose solution is always positive. To motivate it, consider n Ornstein-Uhlenbeck processes $X_i(t)$ given by the SDE's

$$dX_i(t) = -\frac{\alpha}{2}X_i(t) dt + \frac{\sigma}{2}dB_i(t) \quad (21.5)$$

with n uncorrelated Brownian motions B_1, \dots, B_n . Consider the process

$$r(t) := X_1^2(t) + \dots + X_n^2(t) \quad (21.6)$$

then

$$\begin{aligned} dr(t) &= \sum_{i=1}^n 2X_i(t) \left(-\frac{\alpha}{2}X_i(t) dt + \frac{\sigma}{2}dB_i(t) \right) + \frac{1}{2} \sum_{i=1}^n 2\frac{\sigma^2}{4} dt \\ &= -\alpha r(t) dt + \sigma \sum_{i=1}^n X_i(t) dB_i(t) + \frac{n\sigma^2}{4} dt \\ &= \left(\frac{n\sigma^2}{4} - \alpha r(t) \right) dt + \sigma \sqrt{r(t)} \sum_{i=1}^n \frac{X_i(t) dB_i(t)}{\sqrt{r(t)}} \end{aligned} \quad (21.7)$$

The process $W(t) := \sum_{i=1}^n \int_0^t \frac{X_i(s) dB_i(s)}{\sqrt{r(s)}}$ is a continuous martingale with quadratic variation $\langle W(t), W(t) \rangle = \int_0^t ds = t$, hence, a one dimensional Brownian motion. Thus the process $r(t)$ satisfies the SDE

$$dr(t) = \left(\frac{n\sigma^2}{4} - \alpha r(t) \right) dt + \sigma \sqrt{r(t)} dW(t) \quad (21.8)$$

Definition 21.1: A Cox-Ingersoll-Ross process is a process r_t defined by the SDE

$$dr_t = (a - br_t) dt + \sigma \sqrt{r_t} dB_t \quad (21.9)$$

where $a, b, \sigma > 0$ are positive constants.

Theorem 21.2: Let r_t be a Cox-Ingersoll-Ross process starting at $r_0 = x$. Then

a) For $\lambda, \mu > 0$

$$\mathbb{E}\left[e^{\lambda r_T - \mu \int_0^T r_s ds}\right] = e^{-a\phi_{\lambda,\mu}(T) - x\psi_{\lambda,\mu}(T)} \quad (21.10)$$

where

$$\phi_{\lambda,\mu}(t) = -\frac{2}{\sigma^2} \log\left[\frac{2\gamma e^{\frac{t(b+\gamma)}{2}}}{\sigma^2\lambda(e^{\gamma t} - 1) + \gamma - b + e^{\gamma t}(\gamma + b)}\right] \quad (21.11)$$

$$\psi_{\lambda,\mu}(t) = \frac{\lambda(\gamma + b) + e^{\gamma t}(\gamma - b) + 2\mu(e^{\gamma t} - 1)}{\sigma^2\lambda(e^{\gamma t} - 1) + \gamma - b + e^{\gamma t}(\gamma + b)} \quad (21.12)$$

and $\gamma = \sqrt{b^2 + 2\sigma^2\mu}$.

b) For all $r_0 = x > 0$, the following holds: For $a < \sigma^2/2$,

$$\mathbb{P}(\text{there are infinitely many times for which } r_t = 0) = 1. \quad (21.13)$$

For $a \geq \sigma^2/2$,

$$\mathbb{P}(\text{there is at least one time for which } r_t = 0) = 0. \quad (21.14)$$

Proof: Part (a) can be obtained from (16.33), which reduced the left hand side of (21.10) to an expectation for a squared Bessel process, and the corresponding result (20.25). For a direct proof, using martingale methods as in the proof of (20.25), one may look in [4]. Part (b) is proven in exercise 6. ■

Finally we write down a complex version of (21.10) which will be used in the Heston model. It reads as follows.

Corollary 21.3: Let Y_t a Cox-Ingersoll-Ross process given by the SDE

$$dY_t = \kappa(\bar{y} - Y_t)dt + \sigma\sqrt{Y_t}dB_t \quad (21.15)$$

with initial condition $Y_0 = y$ and positive constants $\kappa, \bar{y}, \sigma > 0$. Let $\lambda = a + ib$, $\mu = c + id$. Then

$$\mathbb{E}\left[e^{\frac{\lambda}{2}Y_t - \frac{\mu}{2}\int_0^t Y_s ds}\right] = e^{\frac{\kappa+2f'/f(0)}{\sigma^2}y + \frac{\kappa^2}{\sigma^2}\bar{y}t} e^{\frac{2\kappa\bar{y}}{\sigma^2}\left\{\log\left[\frac{r_f(t)}{r_f(0)}\right] + i(\varphi_f(t) - \varphi_f(0))\right\}} \quad (21.16)$$

where the function f is given by

$$\begin{aligned} f(s) &= \sqrt{\xi} \cosh[\sqrt{\xi}(t-s)] - \gamma \sinh[\sqrt{\xi}(t-s)] \\ &=: r_f(s) e^{i\varphi_f(s)} \end{aligned} \quad (21.17)$$

with a differentiable φ_f and

$$\gamma = \frac{\lambda\sigma^2}{4} - \frac{\kappa}{2} \quad (21.18)$$

$$\xi = \frac{\mu\sigma^2}{4} + \frac{\kappa^2}{4}. \quad (21.19)$$

Proof: The expectation can be reduced to an expectation involving a Bessel process. Recall formula (16.33),

$$\mathbb{E}\left[e^{\frac{\lambda}{2}Y_t - \frac{\mu}{2}\int_0^t Y_s ds}\right] = e^{\frac{\kappa}{\sigma^2}(y + \kappa\bar{y}t)} \mathbb{E}\left[e^{(\frac{\lambda\sigma^2}{8} - \frac{\kappa}{4})X_t - (\frac{\mu\sigma^2}{8} + \frac{\kappa^2}{8})\int_0^t X_s ds}\right] \quad (21.20)$$

where X_t is a squared Bessel process satisfying the SDE

$$dX_t = \frac{4\kappa\bar{y}}{\sigma^2} dt + 2\sqrt{X_t} dx_t \quad (21.21)$$

with initial condition $X_0 = \frac{4y}{\sigma^2} =: x$. Thus, by Corollary 16.3 we obtain

$$\begin{aligned} \mathbb{E}\left[e^{\frac{\lambda}{2}Y_t - \frac{\mu}{2}\int_0^t Y_s ds}\right] &= e^{\frac{\kappa}{\sigma^2}(y + \kappa\bar{y}t)} e^{\frac{f'(0)}{f(0)}\frac{x}{2} + \frac{2\kappa\bar{y}}{\sigma^2}} \left\{ \log\left[\frac{r_f(t)}{r_f(0)}\right] + i(\varphi_f(t) - \varphi_f(0)) \right\} \\ &= e^{\frac{\kappa + 2f'/f(0)}{\sigma^2}y + \frac{\kappa^2}{\sigma^2}\bar{y}t} e^{\frac{2\kappa\bar{y}}{\sigma^2}} \left\{ \log\left[\frac{r_f(t)}{r_f(0)}\right] + i(\varphi_f(t) - \varphi_f(0)) \right\} \end{aligned} \quad (21.22)$$

where the function f is given by

$$f(s) = \sqrt{\xi} \cosh[\sqrt{\xi}(t-s)] - \gamma \sinh[\sqrt{\xi}(t-s)] \quad (21.23)$$

with $\gamma = \frac{\lambda\sigma^2}{4} - \frac{\kappa}{2}$ and $\xi = \frac{\mu\sigma^2}{4} + \frac{\kappa^2}{4}$. ■